

Auditory P300 and N100 components as intermediate phenotypes for psychotic disorder: Familial liability and reliability

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HIGHLIGHTS

- A reliable N100 latency delay was found in unaffected siblings of patients with a psychotic disorder.
- P300 amplitude and latency were not found to be affected in siblings.
- Short-term test–retest reliability of N100 and P300 components were sound across patients, siblings and controls, with the main exception of N100 latency in patients.

ABSTRACT

Objective: Abnormalities of the auditory P300 are a robust finding in patients with psychosis. The purposes of this study were to determine whether patients with a psychotic disorder and their unaffected siblings show abnormalities in P300 and N100 and to establish test–retest reliabilities for these ERP components.

Methods: Using an auditory oddball paradigm, P300 and N100 latency and amplitude were acquired from 19 patients with a psychotic disorder, 28 unaffected siblings, and 37 healthy controls, on two separate occasions. ERP components were compared between groups, using multilevel random regression analyses. Intraclass correlations were used to determine consistency of ERP components between the sessions.

Results: A delayed target N100 latency was found in unaffected siblings. Patients showed significantly delayed P300 latency and diminished P300 amplitude compared to controls. Most ERP parameters showed good test–retest reliability. However, patients did not show sufficient reliability for N100 latency for standard stimuli.

Conclusions: The present study failed to find significant P300 abnormalities in unaffected siblings. However, N100 latency is delayed in siblings and can be reliably measured in all groups for target stimuli, suggesting that this component, rather than P300, may serve as liability marker.

Significance: N100 latency is a promising biomarker for psychosis liability.

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1. Introduction

Schizophrenia is a highly heritable disorder (Gottesman, 1991; Cardno and Gottesman, 2000). However, genetic association studies have, as yet, failed to provide consistent results regarding the precise mode of transmission of the genetic vulnerability. The genetic

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complexity of the disorder has led to the search for intermediate phenotypes with a simpler genetic basis than the dichotomous schizophrenia phenotype in order to facilitate the identification of genetic loci involved in the disorder. Deficits in cognition and information processing, which are prominent in schizophrenia, may be such intermediate phenotype markers. Event-related brain potentials (ERPs) reflect neural activity associated with cognitive information processing (Donchin, 1979) and as such, could be biological intermediate phenotype markers for schizophrenia. The P300 waveform has been considered as a promising candidate intermediate phenotype. P300 amplitude is thought to index

brain activity reflecting attention to incoming stimulus information when representations are updated, as well as attribution of salience to deviant stimuli (Polich, 2007; Turetsky et al., 2007). The P300 latency is considered to be a measure of perceptual processing speed (Polich, 2007). Amplitude reduction of the auditory P300 wave is a robust finding in patients with a psychotic disorder and P300 latency has been shown to be delayed in patients compared to healthy controls (Jeon and Polich, 2001; Bramon et al., 2004). Twin and family studies exhibit moderate to high heritability of the P300 amplitude (Polich and Burns, 1987; O'Connor et al., 1994; Wright et al., 2001; Hall et al., 2006) and, although less consistently, suggest that a significant proportion of the variance in P300 latency may be attributed to genetic factors (Katsanis et al., 1997; Almasy et al., 1999; Wright et al., 2001). Studies investigating the P300 waveform in first-degree relatives of patients with a psychotic disorder have produced mixed results. Some family studies have found reductions in P300 amplitudes in siblings (Kidogami et al., 1991; Schreiber et al., 1992; Frangou et al., 1997; Weisbrod et al., 1999; Kimble et al., 2000; Turetsky et al., 2000), whereas other studies did not find differences between siblings and controls (Karoumi et al., 2000; de Wilde et al., 2008; Sumich et al., 2008). In a meta-analysis, Bramon et al. (2005) showed that siblings displayed normal P300 amplitude, but had a significantly prolonged P300 latency.

Reductions in the amplitude of the earlier auditory N100 evoked potential are also found in patients with a psychotic disorder, reflecting deficits in mechanisms involved in initial sensory processing and early selective attention (Strik et al., 1992; Frangou et al., 1997; Laurent et al., 1999). A study of healthy twins suggested that the reduction in N100 amplitude is highly heritable (Anokhin et al., 2007) and there is some evidence that the reduction in N100 amplitude is also seen in first-degree relatives (Blackwood et al., 1991; Frangou et al., 1997; Turetsky et al., 2008).

Thus, the P300 and N100 waveforms both may be potential intermediate phenotypes since both appear to be abnormal in patients, both appear heritable and there is also evidence, although mixed, that both show abnormalities in healthy first-degree relatives. However, biological markers should not only be meaningfully associated with the disorder and be under significant genetic control; they should also be stable over time to be considered as useful intermediate phenotypes (de Geus, 2002). Studies of healthy control subjects suggest that P300 amplitude has good test–retest reliability (Segalowitz and Barnes, 1993; Mathalon et al., 2000; Walhovd and Fjell, 2002; Winterer et al., 2003) and measurements of N100 amplitude reliability are satisfactorily as well (Segalowitz and Barnes, 1993; Kinoshita et al., 1996; Walhovd and Fjell, 2002; Fuerst et al., 2007). Compared to amplitude, ERP latency generally shows lower test–retest reliability (Polich, 1986; Fabiani et al., 1987; Kinoshita et al., 1996). Relatively few studies have reported short-term (days) test–retest reliability in patients with a psychotic disorder and to our knowledge, no study to date has reported separate test–retest reliabilities for siblings.

The aims of the present study were to investigate P300 and N100 components as potential intermediate phenotypes of psychosis. More specifically, the present study investigated whether these ERP components (i) display abnormalities in patients with a psychotic disorder as well as in their unaffected family members, and (ii) show adequate test–retest reliability across control, sibling and patient groups.

2. Methods

2.1. Subjects

The present study consisted of three groups: (i) 22 patients with a DSM-IV diagnosis of non-affective psychosis (17 schizophrenia, 1

schizophreniform disorder, 2 schizoaffective disorder, 2 psychotic disorder not otherwise specified), (ii) 31 non-psychotic siblings of patients with a non-affective psychosis, and (iii) 39 healthy control participants without a familial history of psychosis. Inclusion criteria were: fluent in Dutch and aged between 18 and 50 years. Subjects were excluded if there was a history of significant head injuries or neurological disorders. All three groups were frequency-matched on age and gender (see Table 1). Patients were recruited from community mental health centres and psychiatric hospitals in the south of the Netherlands and in Belgium. Siblings were recruited through the participating patient, and control participants were recruited from the same geographical regions as the patients through advertising and mailing lists.

Participants were interviewed by trained psychologists using the Comprehensive Assessment of Symptoms and History (CASH; Andreasen et al., 1992). Two verbal subtests (Information and Arithmetic) and two performance subtests (Block Design and Symbol Search) of the Wechsler Adult Intelligence Scale-III (Wechsler, 1997) were used to estimate IQ (Blyler et al., 2000).

The study was approved by the medical ethical committee of Maastricht University Medical Centre. All participants provided written informed consent.

2.2. Experimental procedure

Subjects took part in two identical recording sessions, temporally spaced apart by 11 days ($SD = 7.5$, range: 2–44 days). The oddball paradigm was assessed as part of a larger study, which further included: assessments of cognitive functioning (verbal memory, sustained attention, executive functioning, and processing speed), resting EEG, 40-Hz steady state response, CNV paradigm, three gamma band paradigms, P50 gating paired click-paradigm and MMN oddball paradigm. The order of administration was fixed. The P300 auditory oddball paradigm was administered within the first 15 min of the EEG test session. To assess the presence of psychiatric symptoms at the time of testing, the extended Brief Psychiatric Rating Scale (Lukoff et al., 1986) was administered after the second recording session.

2.3. Auditory oddball paradigm

Within a simple auditory two-tone oddball paradigm, 576 auditory stimuli were presented binaurally by loudspeaker at an approximately 60-dB sound pressure level. Non-targets (1000 Hz tones) and targets (2000 Hz tones) with an immediate rise/fall were presented in a ratio of 7:1 in a pseudo-randomized sequence with a stimulus duration of 50 and a 1000 ms fixed inter-stimulus interval. Subjects were instructed to fixate their eyes on a central cross displayed on a monitor and to press a button in response to targets only.

2.4. ERP recording

Scalp electrode activity was recorded using Neuroscan Synamps and Neuroscan Scan 4.3 software (Neuroscan Inc., Sterling, VA, USA) and was measured at 30 electrode sites of which Fz, Cz and Pz were analysed. Fz, Cz and Pz were chosen for analyses because N100 and P300 responses are largest on the midline locations, probably due to the fact that midline electrodes pick up both left and right hemisphere activity. Furthermore, so far, most studies examining the N100 and P300 component using oddball paradigms, have presented statistical results of midline electrodes only (e.g. Pontifex et al., 2009). Supplementary Figs. S1 and S2 display the ERPs for all electrode sites. These figures show that largest responses were indeed found on the midline. It was therefore decided to present data from Fz, Cz and Pz only. The EEG electrodes

Table 1
Demographic and clinical variables.

	Patients (n = 20)		Siblings (n = 28)		Controls (n = 38)		Statistic (df)	p-Value
	Mean	SD	Mean	SD	Mean	SD		
Age	29.3	6.4	29.2	7.6	28.6	9.1	$F(2.86) = .06$.94
Male sex, n (%)	13 (65%)		17 (61%)		24 (63%)		$\chi^2(2) = .10$.95
Education ^a	5.6	1.7	6.1	1.9	6.0	1.7	$F(2.86) = .53$.59
IQ ^b	97.7	16.2	114.1	17.8	113.4	15.4	$F(2.86) = 7.36$.001
Reaction time	376	83.0	297	48.8	280	40.9	$F(2.86) = 20.37$	<.0001
BPRS total ^c	42.0	13.9	28.3	3.6	26.6	2.1	$F(2.85) = 33.33$	<.0001
<i>Antipsychotic use</i>								
Atypical	n = 12							
Typical	n = 3							

^a Educational level achieved, measured on a nine-point scale from no education (0) to university degree (8).

^b IQ was assessed using a short form of the WAIS-III using Information, Block Design, Digit Symbol and Arithmetic.

^c Brief Psychiatric Rating Scale total scores.

were referenced to the left mastoid. Tin electrodes were used to record bipolarly the vertical (above and below the left eye) and horizontal (at outer canthi of both eyes) electrooculogram (EOG). An electrode at an anterior midline site (AFz) served as ground. Electrode impedances were kept below 5 k Ω . EEG and electrooculogram were digitally filtered with a bandpass of 0.1–100 Hz. Digitization rate and gain were 1000 Hz and 150, respectively, and no notch filter was applied.

2.5. ERP analysis

Analyses were performed using Neuroscan Scan 4.3 (Neuroscan Inc., Sterling, VA, USA). Data were filtered off-line with a 1 Hz high-pass filter (6 dB/oct). Ocular activity was removed using a regression procedure (Gratton et al., 1983). Epochs were segmented at intervals of 100 ms pre-stimulus and 1000 ms post-stimulus. After segmentation, data were filtered using a 30-Hz low-pass filter (6 dB/oct) and data were baseline-corrected. Artefacts were removed both automatically by eliminating epochs that contained signals exceeding $\pm 75 \mu\text{V}$ between 100 ms pre-stimulus and 500 ms post-stimulus, and based on visual inspection. Sweeps with incorrect performance were also rejected, that is target tones to which no response was made and non-targets to which a response was made were excluded from analysis. The ERP components for target and non-target tones were analysed separately.

Components were defined as: N100, the most negative peak occurring between 70 and 160 ms post-stimulus; and P300, the most positive peak between 250 and 500 ms post-stimulus. Latency windows for P300 were defined based on peak latency at Pz. Next, P300 latencies were separately searched for at Fz and Cz, defining them as the peak value within that chosen window. Latency windows for N100 were defined based on peak latency at Cz and were separately searched for at Fz and Pz, defining them as the peak value within the window chosen at Cz. Peak amplitudes were measured from baseline to peak, using the same time windows.

2.6. Statistical analysis

The present data have a hierarchical structure. Multiple observations (level 1) were clustered within subjects (level 2), who were part of families (level 3). Multilevel random regression analysis is the method of choice to deal with data consisting of observations at more than one level in terms of unit of analysis, by taking the variability associated with each level of nesting into account (Snijders and Bosker, 1999). The XT MIXED command in STATA 10.0 (StataCorp, 2007) was used to conduct multilevel linear regression analyses, fitted with maximum likelihood methods. Data from test and retest sessions and three midline electrodes (Fz, Cz, Pz) were entered into the regression analyses. For the standard stimuli, there

was no clear P300 component in a large proportion of the sample. Therefore, we conducted analyses of the P300 components for target stimuli only, whereas the analyses for the N100 components were conducted for the standard as well as target stimuli. Each ERP variable was assessed in a separate analysis. Sex and age were included as covariates for all regression analyses. The group variable was entered as two dummy variables (siblings, patients) comparing associations with the reference group (controls).

Test–retest reliability was estimated by calculating intraclass correlation coefficients (ICCs) for each of the three midline electrodes (Fz, Cz, Pz), carried out using SPSS (SPSS version 16.0 for Windows; SPSS Inc., Chicago). ICCs take both within-subject as well as between-subject variance into account and therefore provide a better measure of test–retest reliability, compared with, for example, Pearson product moment correlations (Bartko, 1991, 1994; Farahat et al., 2003). ICCs were used in a two-way mixed effects model with consistency and single measurements (Farahat et al., 2003). It was decided *a priori* that a reliability coefficient of less than .40 would be considered poor, coefficients between .40–.59 would be considered fair, coefficients between .60–.75 good and coefficients larger than .75 excellent, based on previous accounts of classifying the degree of reliability (Fleiss, 1986; Rentsch et al., 2008).

3. Results

3.1. Sample

Due to technical difficulties, two participants of the patient group and one participant of the control group were excluded. Additionally, one participant of the control group was excluded because of use of antidepressive medication and one sibling was excluded because of use of methylphenidate. One patient's EEG data contained excessive artefacts on both occasions; two control subjects had data containing excessive artefacts on a single occasion. These data were excluded from further analyses. Thus, EEG data from the oddball paradigm were available for 20 patients, 28 siblings and 38 healthy control participants, stemming from 68 different families (including 12 patient–sibling couples, 1 patient–sibling–sibling couple, 2 sibling–sibling couples, 2 control–control couples and 7 single patients, 9 single siblings and 34 single controls). Eighty-three of these subjects had data available from both sessions. Table 1 shows demographic and clinical data.

3.2. ERP data

Descriptive statistics are depicted in Table 2. Grand average waveforms elicited by standard and target tones are presented in Fig. 1. There were no large or significant differences between groups in the number of sweeps analysed.

Table 2
P300 and N100 descriptive statistics.

	Patients (n = 20)		Siblings (n = 28)		Controls (n = 38)	
	Mean	SD	Mean	SD	Mean	SD
<i>N100 latency (ms)</i>						
Standards	127	11.2	127	6.3	124	8.4
Targets	127	12.1	128	9.2	122	11.0
<i>N100 amplitude (μV)</i>						
Standards	-2.9	1.8	-3.6	1.9	-3.9	2.4
Targets	-4.1	2.9	-4.8	2.3	-4.0	2.9
<i>P300 latency (ms)</i>						
Targets	332	46.0	304	39.2	290	38.1
<i>P300 amplitude (μV)</i>						
Targets	7.8	4.9	11.7	4.2	11.0	5.2

3.3. N100

For the standard stimuli, N100 latencies of patients differed significantly from controls ($\beta = .42, p = .048$). Siblings also displayed a significantly longer N100 latency than controls ($\beta = .38, p = .049$). Results were similar for target stimuli: patients ($\beta = .41, p = .04$)

as well as siblings ($\beta = .51, p = .001$) displayed significantly longer N100 latencies compared to controls. As the box plot in Fig. 2 shows, the significant differences between controls and siblings/patients were not produced by outliers.

A trend towards significance was found for the difference in N100 amplitude between patients and controls for standard stimuli ($\beta = .39, p = .09$). There was no significant difference in N100 amplitude between patients and controls for target stimuli ($\beta = -.20, p = .33$), nor were there significant differences in N100 amplitude between siblings and controls (standards: $\beta = .15, p = .49$; targets $\beta = -.23, p = .22$).

3.4. P300

Patients showed a significantly longer P300 latency compared to controls ($\beta = .94, p < .001$). The difference in P300 latency in siblings compared to controls, showed a trend towards significance ($\beta = .32, p = .09$).

Patients showed a significantly reduced P300 amplitude compared to controls ($\beta = -.61, p = .004$), whereas siblings did not differ significantly from controls ($\beta = .15, p = .44$).

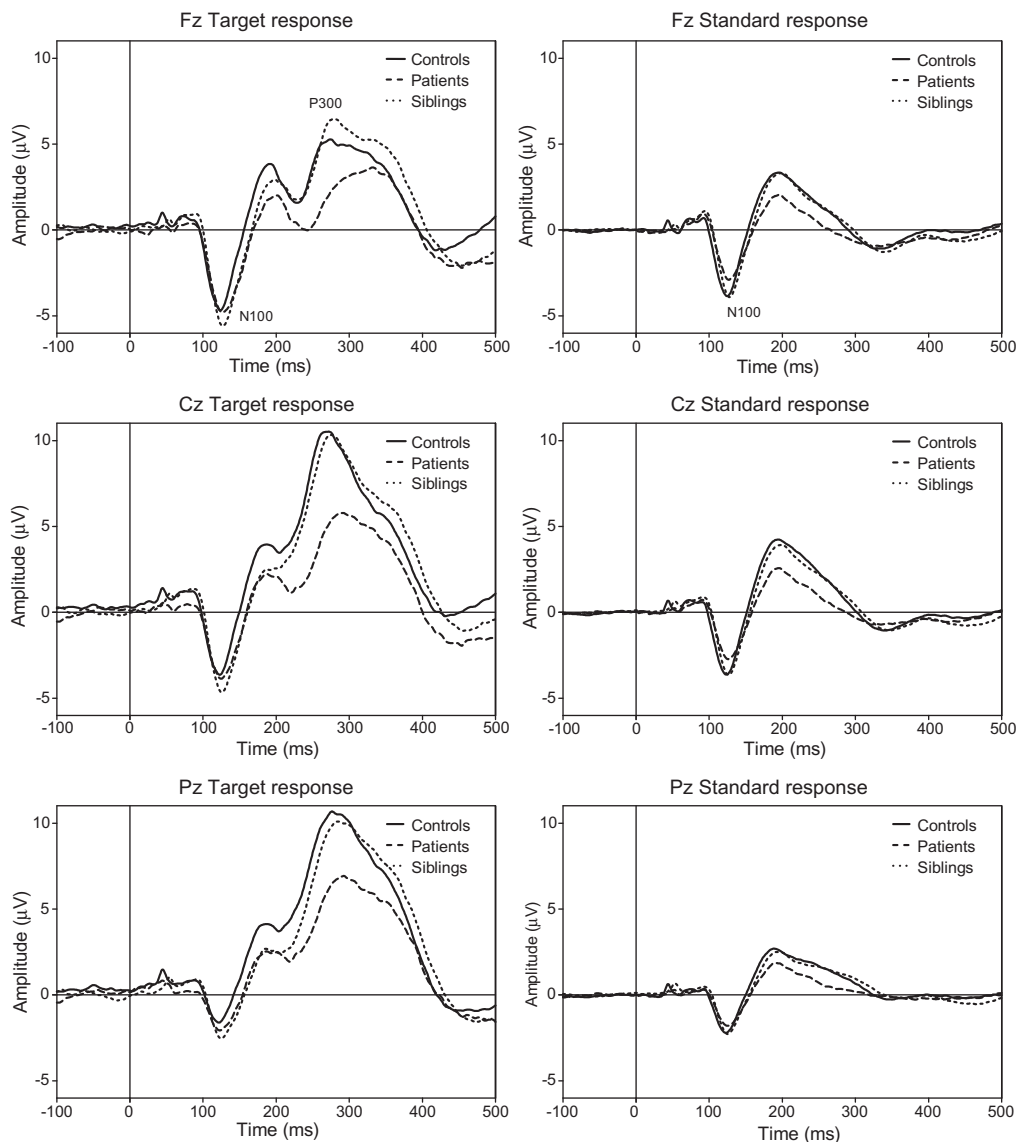


Fig. 1. Grand average event-related potential waveforms elicited by oddball target and standard tones in patients with schizophrenia, unaffected siblings of patients and healthy controls.

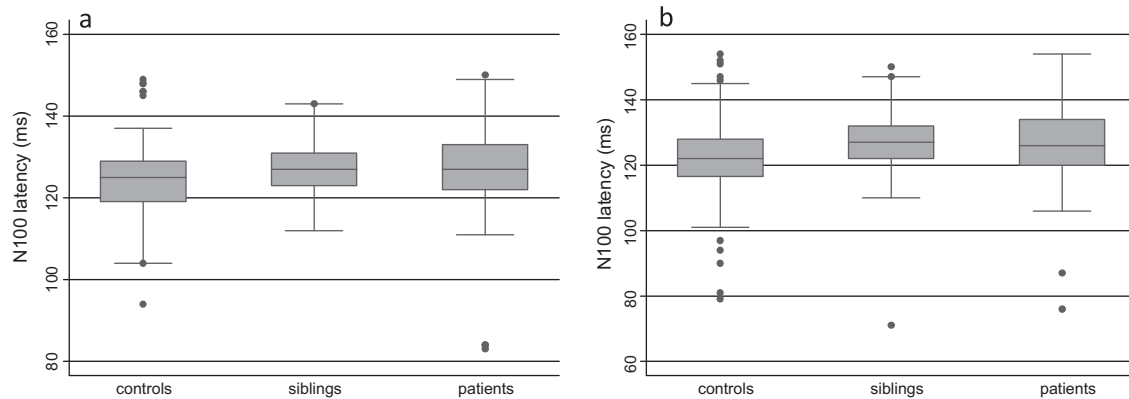


Fig. 2. Box plots for N100 latencies measured at Fz, Cz and Pz, per group for standard stimuli (a) and target stimuli (b).

3.5. Test–retest reliability

The intraclass correlation coefficients as a measure of test–retest reliability are displayed in Table 3. Of the 54 ICCs, 22 were larger than .75 and can therefore be classified as evidence of “excellent” test–retest reliability. Eighteen ICCs could be classified as good, nine as fair and five as poor. Test–retest reliability for all P300 amplitude measures was good to excellent. P300 latency was typically less consistent with all P300 latencies ranging from fair to excellent, with the exception of one. All N100 amplitude measures had ICCs ranging from fair to excellent. ICCs for N100 latency were less consistent; test–retest reliability for the N100 latency parameters for standard stimuli was poor in patients and failed to reach statistical significance.

Table 3
Intraclass correlation coefficients. Two-way mixed effect model with consistency and single measurements.

	Patients ICC (CI)	Siblings ICC (CI)	Controls ICC (CI)
<i>N100 latency standards</i>			
Fz	.20 (–.25–.58)	.86 (.72–.94)**	.85 (.71–.92)**
Cz	.20 (–.27–.59)	.89 (.76–.95)**	.94 (.88–.97)**
Pz	–.07 (–.49–.37)	.61 (.29–.80)**	.72 (.50–.84)**
<i>Targets</i>			
Fz	.74 (.45–.89)**	.63 (.33–.82)**	.73 (.52–.85)**
Cz	.72 (.40–.88)**	.76 (.52–.88)**	.50 (.20–.72)*
Pz	.40 (–.04–.71)*	.45 (.08–.71)*	.34 (.01–.61)*
<i>N100 amplitude standards</i>			
Fz	.92 (.80–.97)**	.92 (.82–.96)**	.77 (.59–.88)**
Cz	.89 (.73–.96)**	.90 (.78–.95)**	.74 (.53–.86)**
Pz	.79 (.54–.91)**	.77 (.55–.89)**	.63 (.37–.79)**
<i>Targets</i>			
Fz	.63 (.26–.84)**	.62 (.31–.81)**	.75 (.55–.87)**
Cz	.76 (.47–.90)**	.61 (.29–.81)**	.76 (.56–.87)**
Pz	.53 (.13–.78)*	.45 (.09–.71)*	.56 (.27–.75)**
<i>P300 latency targets</i>			
Fz	.61 (.24–.83)*	.77 (.56–.89)**	.34 (.01–.61)*
Cz	.64 (.27–.84)*	.76 (.52–.88)**	.61 (.34–.78)**
Pz	.52 (.11–.78)*	.54 (.20–.76)*	.41 (.08–.65)*
<i>P300 amplitude targets</i>			
Fz	.77 (.51–.90)**	.73 (.49–.87)**	.77 (.59–.88)**
Cz	.83 (.61–.93)**	.84 (.67–.93)**	.75 (.55–.87)**
Pz	.85 (.65–.94)**	.80 (.61–.91)**	.74 (.54–.86)**

CI: 95% confidence interval.

* $p < .05$.

** $p < 0.001$.

4. Discussion

4.1. Between-group differences

The present study confirms previous findings of reduced P300 amplitude (Frangou et al., 1997; Weisbrod et al., 1999; Turetsky et al., 2000; Winterer et al., 2003; de Wilde et al., 2008; Sumich et al., 2008) and latency (Bramon et al., 2005) in patients with a psychotic disorder. The unaffected sibling group did not show any significant differences in P300 latency and amplitude, although a trend towards significance was seen for P300 latency. There have been several negative reports regarding P300 abnormalities in unaffected siblings (Karoumi et al., 2000; de Wilde et al., 2008; Sumich et al., 2008). Task difficulty may influence the magnitude of genetic influences on the variance in P300 amplitude (Polich and Burns, 1987; van Beijsterveldt et al., 1998) and may therefore influence the effect size when comparing a genetic risk group with a healthy control group. The oddball task used in the present study is a relatively easy task, which could explain the lack of effect in siblings. However, other studies suggest comparable heritability for easy and difficult tasks (Katsanis et al., 1997; Wright et al., 2001). In their review, Bharath et al. (2000) point out that studies with difficult tasks found P300 differences in high risk groups, whereas relatively simple oddball tasks did not and suggest that high risk studies should use more attention demanding tasks.

There were no significant differences in N100 amplitude between the three groups, in contrast to previous studies finding amplitude reductions in patients (Strik et al., 1992; Laurent et al., 1999) and in first-degree relatives (Blackwood et al., 1991; Frangou et al., 1997). N100 latency, on the other hand, was significantly delayed in patients and in siblings, for non-target as well as target stimuli, thus suggesting that N100 latency may be a manifestation of familial and possibly genetic liability for psychosis. The results are suggestive of slowed processing of auditory information early in the processing sequence, consistent with findings of abnormalities in other early components such as abnormalities in mismatch negativity (Javitt et al., 1993; Alain et al., 1998; Bramon et al., 2004) and impaired auditory sensory gating (Clementz et al., 1997; Boutros et al., 2009). Slowing of initial auditory processing may have downstream consequences for conscious processing of information, as indexed by delayed peaking of later ERP components such as P300.

4.2. Reliability

The reliability of the ERP components recorded with an average interval of 11 days was sound, with the main exception of N100

latencies for standards in patients. P300 amplitude reliability was good to excellent across electrode site and group (ICC .73–.85). The latency of the P300 was less consistent (ICC .34–.76). This is in line with previous studies applying oddball paradigms, showing good test–retest correlation coefficients (r) for P300 amplitude (.5–.8) and good, but generally lower coefficients for P300 latency (.4–.7) in healthy controls (Polich, 1986; Fabiani et al., 1987; Segalowitz and Barnes, 1993).

N100 amplitude reliability ranged from fair to excellent across group and electrode (ICC .45–.92), with somewhat higher reliabilities for standard trials than for target trials. Conform previous findings (Walhovd and Fjell, 2002; Fuerst et al., 2007), N100 amplitude measures were overall more robust than latency measures, in line with the findings for the P300 reliability measures. Previous studies suggest that ERP test–retest reliability tends to follow the topographical distribution of the ERP component, and is greatest where the component is maximal (Walhovd and Fjell, 2002; Williams et al., 2005). N100 latency and amplitude reliabilities did indeed appear lower for Pz than for Fz and Cz. In contrast, the P300, which peaks at Pz, did not show clear signs of this topographical effect.

Given that test–retest reliabilities for the N100 were lower for the Pz in all three subject groups and the N100 ERP is less pronounced at Pz, additional analyses were conducted the most relevant electrode (Cz) only. These analyses resulted in similar effect sizes (β 's) as the three-channel model (Fz, Cz, Pz). The delay in N100 latency remains significant in siblings when only looking at Cz (standards: $\beta = .43$, $p = .04$; targets: $\beta = .45$, $p = .04$), as does the delay in N100 latency for standard stimuli in patients ($\beta = .45$, $p = .047$). The delay in N100 latency for targets just misses significance in patients, due to a reduction in power, but with an effect size similar to the three-channel model ($\beta = .43$, $p = .07$). Thus the analyses restricted to the Cz electrode suggest that the N100 latency delay seen in siblings and patients was not decisively influenced by the Pz electrode.

The usefulness of the P300 component as a biological marker for schizophrenia has been questioned, since it is limited by its lack of specificity, e.g. abnormalities in P300 have been reported in other clinical populations and in family members at risk for other neuropsychiatric disorders, such as alcoholism (Hill et al., 1999), bipolar disorder (Schulze et al., 2008) and Alzheimer's disease (Boutros et al., 1995). Furthermore, oddball paradigms and the electrophysiological methods used in family studies of schizophrenia do not show uniformity, thereby complicating comparisons between studies on test–retest reliability.

4.3. Limitations

Siblings and controls were screened for psychotic disorders and affective disorders and were excluded if they had a lifetime history of psychotic disorder. Six of the siblings in the study sample and six controls met criteria for a lifetime history of major depressive disorder. Since P300 amplitude may show abnormalities in unipolar depression (Gangadhar et al., 1993), including subjects with a diagnosis of major depressive disorder may influence the results. However, in depression, P300 abnormalities appear state-dependent and all siblings and controls with a lifetime history of depression were currently in remission.

The study included patient–sibling pairs as well as single patients and single siblings. It cannot be excluded that the study was biased toward siblings who either do not share the risk gene combinations. Although the significant difference in siblings compared to control subjects for N100 latency suggests that the siblings included in the study may share some of the risk gene combinations, bias toward 'healthier' siblings may explain

the lack of significant findings for P300 abnormalities in the sibling group.

At the time of testing, most patients were on antipsychotic medication, thereby raising the possibility that some of the observed differences between patients and controls may be caused by medication effects. The delayed N100 latency we found in patients was also found in the sibling group, who did not use any psychotropic medication. The delayed latency can therefore not be related to medication status only.

In summary, abnormalities in P300 component are highly consistent in patients with a psychotic disorder. N100 latency is delayed in patients and in siblings of patients with a psychotic disorder and can be reliably measured in siblings, and for target stimuli also in patients, thereby lending support for the use of N100 latency as a biological marker for psychosis liability.

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Appendix A. Supplementary data

Supplementary data associated with this article can be found, in the online version, at [doi:10.1016/j.clinph.2011.02.033](https://doi.org/10.1016/j.clinph.2011.02.033).

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